Probability and Statistics I

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Chapter 5. Distributions of Functions of Random Variables

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Let X be a random variables of the continuous type. If we consider a function of X, say Y = u(X), then Y must also be a random variable that has its own distribution. If we can find its distribution function, say,

$$G(\mathbf{y}) = \mathbb{P}(\mathbf{Y} \le \mathbf{y}) = \mathbb{P}(\mathbf{u}(\mathbf{X}) \le \mathbf{y}),$$

then its pdf is given by $g(\boldsymbol{y}) = G'(\boldsymbol{y}).$

Example 5.1-1 Let X have a gamma distribution with pdf

$$f(\mathbf{x}) = \frac{1}{\Gamma(\alpha)\theta^{\alpha}} \mathbf{x}^{\alpha-1} \mathbf{e}^{-\mathbf{x}/\theta}, \qquad 0 < \mathbf{x} < \infty,$$

where $\alpha > 0, \theta > 0$. Let $Y = e^{\chi}$, so that the support of Y is $1 < y < \infty$. For each y in the support, the distribution function of Y is

$$G(y) = \mathbb{P}(Y \le y) = \mathbb{P}(e^X \le y) = \mathbb{P}(X \le \ln y).$$

That is,

$$G(\mathbf{y}) = \int_0^{\ln \mathbf{y}} \frac{1}{\Gamma(\alpha)\theta^{\alpha}} \mathbf{x}^{\alpha-1} \mathbf{e}^{-\mathbf{x}/\theta} d\mathbf{x}$$

and thus the pdf g(y) = G'(y) of Y is

$$g(y) = rac{1}{\Gamma(\alpha) heta^{lpha}} (\ln y)^{lpha - 1} e^{-(\ln y)/ heta} \left(rac{1}{y}
ight), \qquad 1 < y < \infty$$

Equivalently, we have

$$g(\mathbf{y}) = rac{1}{\Gamma(\alpha)\theta^{lpha}} rac{(\ln \mathbf{y})^{lpha - 1}}{\mathbf{y}^{1 + 1/ heta}} \qquad 1 < \mathbf{y} < \infty.$$

which is called a *loggamma* pdf.

Determination of g(y) from f(x)

Let X be a continuous r.v. with pdf f(x). If the transformation y = u(x) is one-to-one and has the inverse transformation

$$\boldsymbol{x} = \boldsymbol{u}^{-1}(\boldsymbol{y}) = \boldsymbol{v}(\boldsymbol{y})$$

then the pdf of Y is given by

$$g(\mathbf{y}) = f(\mathbf{v}(\mathbf{y})) |\mathbf{v}'(\mathbf{y})|, \quad \mathbf{y} \in \mathcal{S}_{\mathbf{y}}$$

where S_{Y} is the support of Y.

Example 5.1-2 Let Y = 2X + 3. Find the pdf of Y if X is a uniform r.v. over (-1, 2). Ans: $g(y) = \begin{cases} \frac{1}{6}, & \text{if } 1 < y < 7, \\ 0, & \text{otherwise.} \end{cases}$ Example 5.1-3 Let X have the pdf $f(x) = xe^{-x^2/2}, 0 < x < \infty$. Find the pdf of $Y = X^2$. Ans: $g(y) = \cdots$ Here are some examples when the transformation Y = u(X) is not one-to-one.

Example 5.1-4 Let $Y = X^2$. Find the pdf of Y when the distribution of X is N(0, 1). Ans. $g(y) = \frac{1}{\sqrt{2\pi y}} \exp(-y/2), \quad 0 < y < \infty$. Example 5.1-5 Let $Y = X^2$. Find the pdf of Y when the distribution of X is

$$f(\mathbf{x}) = \frac{\mathbf{x}^2}{3}, \qquad -1 < \mathbf{x} < 2.$$

Ans.
$$g(y) = \begin{cases} \frac{\sqrt{y}}{3}, & \text{if } 0 < y < 1\\ \frac{\sqrt{y}}{6}, & \text{if } 1 < y < 4.. \end{cases}$$

Discuss Theorem 5.1-1 about the simulation of r.v.'s.

Exercises from textbook: Section 5.1: 1, 3, 4ab, 5, 10, 11, 15