

# Math 362: Mathematical Statistics II

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# Chapter 7. Inference Based on The Normal Distribution

§ 7.1 Introduction

§ 7.2 Comparing  $\frac{\bar{Y}-\mu}{\sigma/\sqrt{n}}$  and  $\frac{\bar{Y}-\mu}{S/\sqrt{n}}$

§ 7.3 Deriving the Distribution of  $\frac{\bar{Y}-\mu}{S/\sqrt{n}}$

§ 7.4 Drawing Inferences About  $\mu$

§ 7.5 Drawing Inferences About  $\sigma^2$

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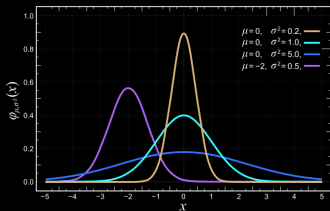
**Carl Friedrich Gauss**   
discovered the normal distribution in 1809 as a way to rationalize the **method of least squares**.

(1777-1855)



**Marquis de Laplace** proved the **central limit theorem** in 1810, consolidating the importance of the normal distribution in statistics. 

(1749-1827)



<b>Notation</b>	$\mathcal{N}(\mu, \sigma^2)$
<b>Parameters</b>	$\mu \in \mathbb{R}$ = mean ( <b>location</b> ) $\sigma^2 > 0$ = variance ( <b>squared scale</b> )
<b>Support</b>	$x \in \mathbb{R}$
<b>PDF</b>	$\frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2}$
<b>CDF</b>	$\frac{1}{2} \left[ 1 + \operatorname{erf}\left(\frac{x-\mu}{\sigma\sqrt{2}}\right) \right]$
<b>Quantile</b>	$\mu + \sigma\sqrt{2} \operatorname{erf}^{-1}(2p - 1)$
<b>Mean</b>	$\mu$
<b>Median</b>	$\mu$
<b>Mode</b>	$\mu$
<b>Variance</b>	$\sigma^2$
<b>MAD</b>	$\sigma\sqrt{2/\pi}$
<b>Skewness</b>	0
<b>Ex. kurtosis</b>	0
<b>Entropy</b>	$\frac{1}{2} \log(2\pi e\sigma^2)$
<b>MGF</b>	$\exp(\mu t + \sigma^2 t^2 / 2)$
<b>CF</b>	$\exp(i\mu t - \sigma^2 t^2 / 2)$
<b>Fisher information</b>	$\mathcal{I}(\mu, \sigma) = \begin{pmatrix} 1/\sigma^2 & 0 \\ 0 & 2/\sigma^2 \end{pmatrix}$ $\mathcal{I}(\mu, \sigma^2) = \begin{pmatrix} 1/\sigma^2 & 0 \\ 0 & 1/(2\sigma^4) \end{pmatrix}$
<b>Kullback-Leibler divergence</b>	$D_{\text{KL}}(\mathcal{N}_0 \parallel \mathcal{N}_1) = \frac{1}{2} \left\{ \left( \frac{\sigma_0}{\sigma_1} \right)^2 + \frac{(\mu_1 - \mu_0)^2}{\sigma_1^2} - 1 + 2 \ln \frac{\sigma_1}{\sigma_0} \right\}$

[https://en.wikipedia.org/wiki/Normal\\_distribution](https://en.wikipedia.org/wiki/Normal_distribution)

## Test for normal parameters (one sample test)

Let  $Y_1, \dots, Y_n$  be a random sample from  $N(\mu, \sigma^2)$ .

**Prob. 1** Find a test statistic  $\Lambda$  in order to test  $H_0 : \mu = \mu_0$  v.s.  $H_1 : \mu \neq \mu_0$ .

When  $\sigma^2$  is known: 
$$\Lambda = \frac{\bar{Y} - \mu_0}{\sigma/\sqrt{n}} \sim N(0, 1)$$

When  $\sigma^2$  is unknown: 
$$\Lambda = ? \quad \Lambda \stackrel{?}{=} \frac{\bar{Y} - \mu_0}{s/\sqrt{n}} \sim ?$$

**Prob. 2** Find a test statistic  $\Lambda$  in order to test  $H_0 : \sigma^2 = \sigma_0^2$  v.s.  $H_1 : \sigma^2 \neq \sigma_0^2$ .

**Prob. 1** Find a test statistic for  $H_0 : \mu = \mu_0$  v.s.  $H_1 : \mu \neq \mu_0$ , with  $\sigma^2$  unknown

**Sol.** Composite-vs-composite test with:

$$\omega = \{(\mu, \sigma^2) : \mu = \mu_0, \sigma^2 > 0\}$$

$$\Omega = \{(\mu, \sigma^2) : \mu \in \mathbb{R}, \sigma^2 > 0\}$$

The MLE under the two spaces are:

$$\omega_{\theta} = (\mu_{\theta}, \sigma_{\theta}^2) : \quad \mu_{\theta} = \mu_0 \quad \text{and} \quad \sigma_{\theta}^2 = \frac{1}{n} \sum_{i=1}^n (y_i - \mu_0)^2 \quad (\text{Under } \omega)$$

$$\Omega_{\theta} = (\mu_{\theta}, \sigma_{\theta}^2) : \quad \mu_{\theta} = \bar{y} \quad \text{and} \quad \sigma_{\theta}^2 = \frac{1}{n} \sum_{i=1}^n (y_i - \bar{y})^2 \quad (\text{Under } \Omega)$$

$$L(\mu, \sigma^2) = (2\pi\sigma^2)^{-n} \exp\left(-\frac{1}{2} \sum_{i=1}^n \left(\frac{y_i - \mu}{\sigma}\right)^2\right)$$

$$L(\omega_e) = \dots = \left[ \frac{ne^{-1}}{2\pi \sum_{i=1}^n (y_i - \mu_0)^2} \right]^{n/2}$$

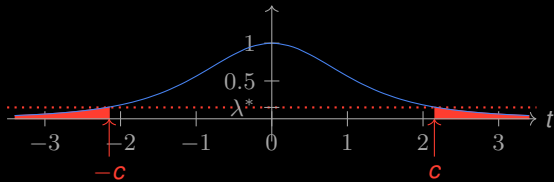
$$L(\Omega_e) = \dots = \left[ \frac{ne^{-1}}{2\pi \sum_{i=1}^n (y_i - \bar{y})^2} \right]^{n/2}$$



Hence,

$$\begin{aligned}\lambda &= \frac{L(\omega_e)}{L(\Omega_e)} = \left[ \frac{\sum_{i=1}^n (y_i - \bar{y})^2}{\sum_{i=1}^n (y_i - \mu_0)^2} \right]^{n/2} = \dots = \left[ 1 + \frac{n(\bar{y} - \mu_0)^2}{\sum_{i=1}^n (y_i - \bar{y})^2} \right]^{-n/2} \\ &= \left[ 1 + \frac{1}{n-1} \left( \frac{\bar{y} - \mu_0}{\sqrt{\frac{1}{n-1} \sum_{i=1}^n (y_i - \bar{y})^2} / \sqrt{n}}} \right)^2 \right]^{-n/2} \\ &= \left[ 1 + \frac{1}{n-1} \left( \frac{\bar{y} - \mu_0}{s / \sqrt{n}} \right)^2 \right]^{-n/2} \\ &= \left[ 1 + \frac{t^2}{n-1} \right]^{-n/2}, \quad t = \frac{\bar{y} - \mu_0}{s / \sqrt{n}}\end{aligned}$$

$$\lambda(t) = \left(1 + \frac{t^2}{n-1}\right)^{-\frac{n}{2}}$$



$$\lambda \in (0, \lambda^*] \quad \Leftrightarrow \quad |t| \geq c.$$

Finally, the test statistic is

$$T = \frac{\bar{Y} - \mu_0}{S/\sqrt{n}}$$

$$\text{with } \bar{Y} = \frac{1}{n} \sum_{i=1}^n Y_i \quad \text{and} \quad S^2 = \frac{1}{n-1} \sum_{i=1}^n (Y_i - \bar{Y})^2.$$

The critical region takes the form:  $|t| \geq c$ .

**Question:** Find the exact distribution of  $T$ .

**Prob. 2** Find a test statistic for  $H_0 : \sigma^2 = \sigma_0^2$  v.s.  $H_1 : \sigma^2 \neq \sigma_0^2$ , with  $\mu$  unknown

**Sol.** Composite-vs-composite test with:

$$\omega = \{(\mu, \sigma^2) : \mu \in \mathbb{R}, \sigma^2 = \sigma_0^2\}$$

$$\Omega = \{(\mu, \sigma^2) : \mu \in \mathbb{R}, \sigma^2 > 0\}$$

The MLE under the two spaces are:

$$\omega_{\theta} = (\mu_{\theta}, \sigma_{\theta}^2) : \quad \mu_{\theta} = \bar{y} \quad \text{and} \quad \sigma_{\theta}^2 = \sigma_0^2 \quad (\text{Under } \omega)$$

$$\Omega_{\theta} = (\mu_{\theta}, \sigma_{\theta}^2) : \quad \mu_{\theta} = \bar{y} \quad \text{and} \quad \sigma_{\theta}^2 = \frac{1}{n} \sum_{i=1}^n (y_i - \bar{y})^2 \quad (\text{Under } \Omega)$$

$$L(\mu, \sigma^2) = (2\pi\sigma^2)^{-n} \exp\left(-\frac{1}{2} \sum_{i=1}^n \left(\frac{y_i - \mu}{\sigma}\right)^2\right)$$

$$L(\omega_e) = (2\pi\sigma^2)^{-n} \exp\left(-\frac{1}{2} \sum_{i=1}^n \left(\frac{y_i - \bar{y}}{\sigma_0}\right)^2\right)$$

$$L(\Omega_e) = \dots = \left[ \frac{ne^{-1}}{2\pi \sum_{i=1}^n (y_i - \bar{y})^2} \right]^{n/2}$$

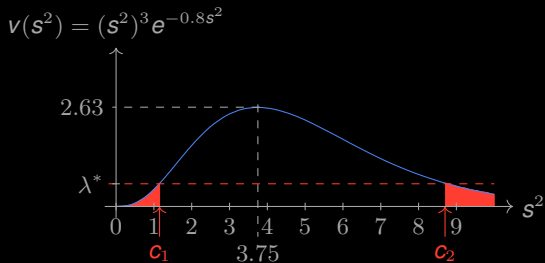
Hence,

$$\begin{aligned}\lambda &= \frac{L(\omega_e)}{L(\Omega_e)} = \left[ \frac{\sum_{i=1}^n (y_i - \bar{y})^2}{n\sigma_0^2} \right]^{n/2} \exp \left( -\frac{1}{2} \sum_{i=1}^n \left( \frac{y_i - \bar{y}}{\sigma_0} \right)^2 + \frac{n}{2} \right) \\ &= \left[ \frac{\frac{1}{n-1} \sum_{i=1}^n (y_i - \bar{y})^2}{\frac{n}{n-1} \sigma_0^2} \right]^{n/2} \exp \left( -\frac{n-1}{2\sigma_0^2} \frac{1}{n-1} \sum_{i=1}^n (y_i - \bar{y})^2 + \frac{n}{2} \right) \\ &= \left[ \frac{\mathbf{s}^2}{\frac{n}{n-1} \sigma_0^2} \right]^{n/2} \exp \left( -\frac{n-1}{2\sigma_0^2} \mathbf{s}^2 + \frac{n}{2} \right)\end{aligned}$$

⇓

$$\lambda(\mathbf{s}^2) = \left[ \frac{\mathbf{s}^2}{\frac{n}{n-1} \sigma_0^2} \right]^{n/2} \exp \left( -\frac{n-1}{2\sigma_0^2} \mathbf{s}^2 + \frac{n}{2} \right) \iff \nu(\mathbf{s}^2) = (\mathbf{s}^2)^{\frac{n}{2}} e^{-\lambda \mathbf{s}^2}$$

By setting  $n = 6$  and  $\lambda = 0.8$ , we see ...



This suggests that the critical region should be of the form in terms of  $s^2$ :

$$(0, c_1) \cup (c_2, \infty)$$

For convenience, we put  $\alpha/2$  mass on each tails of  $S^2$ :

Find  $c_1$  and  $c_2$  such that

$$\int_0^{c_1} f_{S^2}(z) dz = \int_{c_2}^{\infty} f_{S^2}(z) dz = \frac{\alpha}{2}.$$

Finally, the test statistic is

$$\boxed{S^2 = \frac{1}{n-1} \sum_{i=1}^n (Y_i - \bar{Y})^2} \quad \text{with} \quad \bar{Y} = \frac{1}{n} \sum_{i=1}^n Y_i$$

**Question:** Find the exact distribution of  $S^2$ .