

Math 362: Mathematical Statistics II

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Chapter 7. Inference Based on The Normal Distribution

§ 7.1 Introduction

§ 7.2 Comparing $\frac{\bar{Y}-\mu}{\sigma/\sqrt{n}}$ and $\frac{\bar{Y}-\mu}{S/\sqrt{n}}$

§ 7.3 Deriving the Distribution of $\frac{\bar{Y}-\mu}{S/\sqrt{n}}$

§ 7.4 Drawing Inferences About μ

§ 7.5 Drawing Inferences About σ^2

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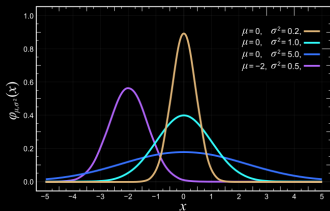
Carl Friedrich Gauss 
discovered the normal distribution in 1809 as a way to rationalize the **method of least squares**.

(1777-1855)



Marquis de Laplace proved the **central limit theorem**  in 1810, consolidating the importance of the normal distribution in statistics.

(1749-1827)



Notation	$\mathcal{N}(\mu, \sigma^2)$
Parameters	$\mu \in \mathbb{R}$ = mean (location) $\sigma^2 > 0$ = variance (squared scale)
Support	$x \in \mathbb{R}$
PDF	$\frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2}$
CDF	$\frac{1}{2} \left[1 + \operatorname{erf}\left(\frac{x-\mu}{\sigma\sqrt{2}}\right) \right]$
Quantile	$\mu + \sigma\sqrt{2} \operatorname{erf}^{-1}(2p - 1)$
Mean	μ
Median	μ
Mode	μ
Variance	σ^2
MAD	$\sigma\sqrt{2/\pi}$
Skewness	0
Ex. kurtosis	0
Entropy	$\frac{1}{2} \log(2\pi e\sigma^2)$
MGF	$\exp(\mu t + \sigma^2 t^2 / 2)$
CF	$\exp(i\mu t - \sigma^2 t^2 / 2)$
Fisher information	$\mathcal{I}(\mu, \sigma) = \begin{pmatrix} 1/\sigma^2 & 0 \\ 0 & 2/\sigma^2 \end{pmatrix}$ $\mathcal{I}(\mu, \sigma^2) = \begin{pmatrix} 1/\sigma^2 & 0 \\ 0 & 1/(2\sigma^4) \end{pmatrix}$
Kullback-Leibler divergence	$D_{\text{KL}}(\mathcal{N}_0 \parallel \mathcal{N}_1) = \frac{1}{2} \left\{ \left(\frac{\sigma_0}{\sigma_1} \right)^2 + \frac{(\mu_1 - \mu_0)^2}{\sigma_1^2} - 1 + 2 \ln \frac{\sigma_1}{\sigma_0} \right\}$

https://en.wikipedia.org/wiki/Normal_distribution

Test for normal parameters (one sample test)

Let Y_1, \dots, Y_n be a random sample from $N(\mu, \sigma^2)$.

Prob. 1 Find a test statistic Λ in order to test $H_0 : \mu = \mu_0$ v.s. $H_1 : \mu \neq \mu_0$.

When σ^2 is known: $\Lambda = \frac{\bar{Y} - \mu_0}{\sigma/\sqrt{n}} \sim N(0, 1)$

When σ^2 is unknown: $\Lambda = ?$

Prob. 2 Find a test statistic Λ in order to test $H_0 : \mu = \mu_0$ v.s. $H_1 : \mu > \mu_0$.

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Prob. 1 Find a test statistic for $H_0 : \mu = \mu_0$ v.s. $H_1 : \mu \neq \mu_0$, with σ^2 unknown

Sol. Composite-vs-composite test with:

$$\omega = \{(\mu, \sigma^2) : \mu = \mu_0, \sigma^2 > 0\}$$

$$\Omega = \{(\mu, \sigma^2) : \mu \in \mathbb{R}, \sigma^2 > 0\}$$

The MLE under the two spaces are:

$$\omega_{\theta} = (\mu_{\theta}, \sigma_{\theta}^2) : \quad \mu_{\theta} = \mu_0 \quad \text{and} \quad \sigma_{\theta}^2 = \frac{1}{n} \sum_{i=1}^n (y_i - \mu_0)^2 \quad (\text{Under } \omega)$$

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$$L(\mu, \sigma^2) = (2\pi\sigma^2)^{-n} \exp\left(-\frac{1}{2} \sum_{i=1}^n \left(\frac{y_i - \mu}{\sigma}\right)^2\right)$$

$$L(\omega_\theta) = \dots = \left[\frac{ne^{-1}}{2\pi \sum_{i=1}^n (y_i - \mu_0)^2} \right]^{n/2}$$

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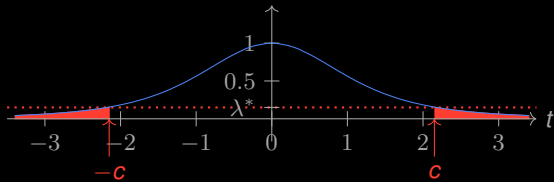
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Hence,

$$\begin{aligned}\lambda &= \frac{L(\omega_e)}{L(\Omega_e)} = \left[\frac{\sum_{i=1}^n (y_i - \bar{y})^2}{\sum_{i=1}^n (y_i - \mu_0)^2} \right]^{n/2} = \dots = \left[1 + \frac{n(\bar{y} - \mu_0)^2}{\sum_{i=1}^n (y_i - \bar{y})^2} \right]^{-n/2} \\ &= \left[1 + \frac{1}{n-1} \left(\frac{\bar{y} - \mu_0}{\sqrt{\frac{1}{n-1} \sum_{i=1}^n (y_i - \bar{y})^2} / \sqrt{n}}} \right)^2 \right]^{-n/2} \\ &= \left[1 + \frac{1}{n-1} \left(\frac{\bar{y} - \mu_0}{s / \sqrt{n}} \right)^2 \right]^{-n/2} \\ &= \left[1 + \frac{t^2}{n-1} \right]^{-n/2}, \quad t = \frac{\bar{y} - \mu_0}{s / \sqrt{n}}\end{aligned}$$

$$\lambda(t) = \left(1 + \frac{t^2}{n-1}\right)^{-\frac{n}{2}}$$



$$\lambda \in (0, \lambda^*] \quad \Leftrightarrow \quad |t| \geq c.$$

Finally, the test statistic is

$$T = \frac{\bar{Y} - \mu_0}{S/\sqrt{n}}$$

$$\text{with } \bar{Y} = \frac{1}{n} \sum_{i=1}^n Y_i \quad \text{and} \quad S^2 = \frac{1}{n-1} \sum_{i=1}^n (Y_i - \bar{Y})^2.$$

The critical region takes the form: $|t| \geq c$.

Question: Find the exact distribution of T .

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Hence,

$$\begin{aligned}\lambda &= \frac{L(\omega_e)}{L(\Omega_e)} = \left[\frac{\sum_{i=1}^n (y_i - \bar{y})^2}{n\sigma_0^2} \right]^{n/2} \exp \left(-\frac{1}{2} \sum_{i=1}^n \left(\frac{y_i - \bar{y}}{\sigma_0} \right)^2 + \frac{n}{2} \right) \\ &= \left[\frac{\frac{1}{n-1} \sum_{i=1}^n (y_i - \bar{y})^2}{\frac{n}{n-1} \sigma_0^2} \right]^{n/2} \exp \left(-\frac{n-1}{2\sigma_0^2} \frac{1}{n-1} \sum_{i=1}^n (y_i - \bar{y})^2 + \frac{n}{2} \right) \\ &= \left[\frac{s^2}{\frac{n}{n-1} \sigma_0^2} \right]^{n/2} \exp \left(-\frac{n-1}{2\sigma_0^2} s^2 + \frac{n}{2} \right)\end{aligned}$$

↓

$$\lambda(s^2) = \left[\frac{s^2}{\frac{n}{n-1} \sigma_0^2} \right]^{n/2} \exp \left(-\frac{n-1}{2\sigma_0^2} s^2 + \frac{n}{2} \right) \iff v(s^2) = (s^2)^{\frac{n}{2}} e^{-\lambda s^2}$$

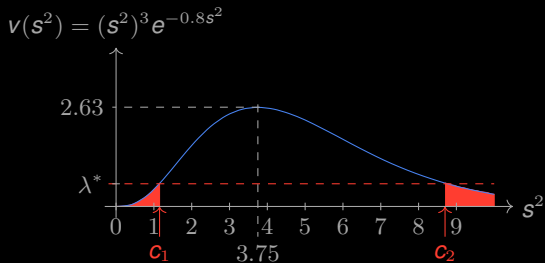
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⇓

$$\lambda(\mathbf{s}^2) = \left[\frac{\mathbf{s}^2}{\frac{n}{n-1} \sigma_0^2} \right]^{n/2} \exp \left(-\frac{n-1}{2\sigma_0^2} \mathbf{s}^2 + \frac{n}{2} \right) \iff \nu(\mathbf{s}^2) = (\mathbf{s}^2)^{\frac{n}{2}} e^{-\lambda \mathbf{s}^2}$$

By setting $n = 6$ and $\lambda = 0.8$, we see ...



This suggests that the critical region should be of the form in terms of s^2 :

$$(0, c_1) \cup (c_2, \infty)$$

For convenience, we put $\alpha/2$ mass on each tails of S^2 :

Find c_1 and c_2 such that

$$\int_0^{c_1} f_{S^2}(z) dz = \int_{c_2}^{\infty} f_{S^2}(z) dz = \frac{\alpha}{2}.$$

Finally, the test statistic is

$$\boxed{S^2 = \frac{1}{n-1} \sum_{i=1}^n (Y_i - \bar{Y})^2} \quad \text{with} \quad \bar{Y} = \frac{1}{n} \sum_{i=1}^n Y_i$$

Question: Find the exact distribution of S^2 .

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